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Coronavirus disease-19 is associated with decreased treatment access and worsened outcomes in malignant brain tumor patients

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ABSTRACT

**Background:** The global coronavirus disease-19 (COVID-19) pandemic has resulted in procedural delays around the world; however, timely and aggressive surgical resection for malignant brain tumor patients is essential for outcome optimization. To investigate the association between COVID-19 and outcomes of these patients, we queried the 2020 National Inpatient Sample (NIS) for differences in rates of surgical resection, time to surgery, mortality, and discharge disposition between patients with and without confirmed COVID-19 infection.

**Methods:** Patient data were taken from the NIS from April 2020 to December 2020. COVID-19 diagnosis was determined with the International Classification of Diseases, Tenth Revision, Clinical Modification code U07.1.

**Results:** A total of 30,671 malignant brain tumor patients met inclusion criteria and 738 (2.4%) patients had a confirmed COVID-19 diagnosis. COVID-19-positive patients had lower likelihood of receiving surgery (Odds ratio [OR] 0.43, 95% confidence interval [CI] 0.29–0.63, P < 0.0001), increased likelihood of mortality (OR 2.18, 95% CI 1.78–2.66, P < 0.0001), and increased likelihood of non-routine discharge (OR 1.25, 95% CI 1.13–1.39, P < 0.0001). Notably, COVID patients receiving surgery were not associated with surgical delay (P = 0.17).

**Conclusion:** COVID-19 infection was associated with worse patient outcome in malignant brain tumor patients, including decreased likelihood of receiving surgery, increased likelihood of mortality, and increased likelihood of non-routine discharge. Our study highlights the need to balance the risks and benefits of delaying surgery for malignant brain tumor patients with COVID-19. Although the COVID-19 pandemic is no longer a public health emergency, understanding the pandemic's impact on outcome provides important insight in effective triage for these patients in the situations where resources are limited.

Keywords: Brain tumor, Coronavirus, Coronavirus disease-19, Malignant, National inpatient sample

# INTRODUCTION

SARS-CoV-2, the virus that causes coronavirus disease-19 (COVID-19), was first identified in December 2019 and by March 2020, it was declared a global pandemic.<sup>[36]</sup> COVID-19 strained the health-care system resulting in the triaging of patients and rationing of supplies.<sup>[2]</sup> In response, a

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series of rapidly evolving guidelines for healthcare were created to minimize the spread of COVID-19, ensure patient safety, and divert resources to the care of COVID-19 patients. As a result, the management of many chronic diseases, including cancer, was drastically altered.<sup>[1,11]</sup> Hospitals reduced or stopped elective surgeries following the guidance of state and federal agencies and medical associations.<sup>[21,22]</sup> In this instance, the definition of "elective" included a number of cancer cases and, as such, many neurosurgical departments suffered delays or cancellations in malignant tumor resections.<sup>[5,12]</sup> By the end of March 2020, an estimated 18-44.8% of all cancer surgeries were canceled in the United States.<sup>[7]</sup> In turn, this led to the American College of Surgeons and the American Association of Neurological Surgeons classifying malignant brain tumors as requiring emergent intervention due the poor prognosis of the tumors and worse outcomes if cancer surgeries were delayed by 3 months.<sup>[8,26,32]</sup> Still, neurosurgeons had to balance the benefits of emergently operating on COVID-19-positive patients with malignant tumors and the unknown risk of COVID-19 on outcomes.

Approximately 25,000 cases of malignant central nervous system cancers occur annually in the United States, representing only 1.4% of total cancer incidence.[30,31] Metastases are the most common malignant brain tumor while the most common primary malignant brain tumor is glioblastoma, which accounts for 14.2% of all brain tumors. Malignant brain tumors as a whole have a very poor prognosis (35.7% 5-year relative survival rate).<sup>[23]</sup> The current standard of practice to improve patient survival is surgical resection of high-grade brain tumors with subsequent chemoradiation.<sup>[20]</sup> Timely and aggressive surgical resection of malignant brain tumor increases overall survival in patients.<sup>[13,18,27]</sup> Resection of the tumor not only reduces the tumor burden but also relieves the mass effect on the brain to preserve neurological function.<sup>[35]</sup> However, resection is not curative as malignant brain tumors have high reoccurrence rates.<sup>[37]</sup> As such, given both the poor prognosis and rapid expansion of such tumors, prompt surgical intervention is vital to improving patient outcomes and survival.

For patients with malignant brain tumors who are COVID-19-positive, the risks and benefits of urgent resection must be weighed. The American Society of Anesthesiologists guidelines recommend waiting 7 weeks after COVID-19 diagnosis for surgeries when possible to minimize postoperative risks.<sup>[3]</sup> However, as malignant tumors are often considered emergent, waiting 7 weeks is not always an option. Notably, we have demonstrated in national data that COVID-19 diagnosis has been associated with delays in benign brain tumor surgery. However, the effect of COVID diagnosis on malignant tumors with the described more urgent indications remains unclear. Further, the full risks associated with tumor resection for COVID-19 patients

are not known due to limited research and data; prior limited studies have shown an association between COVID-19 and malignant brain tumors with poorer outcomes but are limited by small sample size.<sup>[16]</sup> As the pandemic transitions into an endemic phase, it is vital to understand the effect of COVID-19 on postoperative outcomes to maximize treatment efficiency and patient survival, and to optimize further triage management of these patients in future public health emergencies.

In this study, we utilize the National Inpatient Sample (NIS) database to investigate the national rates of surgical resection, mortality, and unfavorable discharge in COVID-19-positive patients with malignant brain tumors. To the best of our knowledge, our study is the first to use national data to analyze the association of a COVID-19 diagnosis with the outcomes of likelihood of receiving surgery, mortality, and discharge disposition in malignant brain tumor patients.

# MATERIALS AND METHODS

## Data source

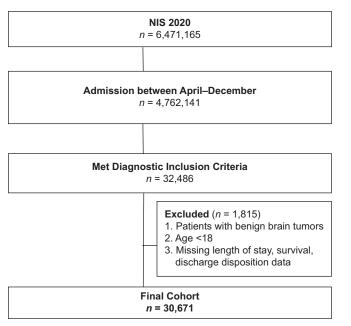
The Healthcare Cost and Utilization Project (HCUP), an initiative by the Agency for Healthcare Research and Quality (AHRQ), releases the NIS annually. This comprehensive database is the most extensive publicly available inpatient health-care database in the United States and covers over 97% of the national population. The NIS comprises approximately 7 million hospital stays each year without weighting, and about 35 million hospitalizations with weighting.<sup>[19]</sup>

# **Study population**

This study is a retrospective database analysis of NIS data from April 2020 to December 2020, which are the first dates to include the International Classification of Diseases, Tenth Revision (ICD-10) diagnosis code (U07.1) for COVID-19 infection. Patients diagnosed with malignant neoplasms of the brain were included in the study. Patients with benign brain tumors were excluded from the cohort. A full list of ICD-10 codes used in inclusion/exclusion criteria can be found in Supplementary Table S1. Patients who were under 18 years old, missing in-hospital survival data, missing length of stay data, or missing discharge disposition were also excluded from the study. Figure 1 provides a breakdown of the patient selection process.

# Study variables

To conduct univariate and multivariate analyses, data were extracted from the NIS database for patient, hospital, and clinical information. COVID-19 was the primary exposure variable (ICD-10 code U07.1). Patient characteristics analyzed included age, sex, race, insurance status/payer, All



**Figure 1:** Flowchart detailing the inclusion and exclusion process in determining the final study cohort. n = Number, NIS = National Inpatient Sample.

Patient Refined Diagnosis Related Groups (APR-DRG) Risk of, APR-DRG Severity of illness, Elixhauser comorbidity index (ECI), patient residential location, and income. Hospital characteristics include bed size (i.e., number of hospital beds), control/ownership of hospital, census division of the hospital, and census region of hospital. Outcome variables included surgical resection (on a binary yes/no scale), in-hospital mortality, discharge disposition (unfavorable vs. favorable), and transfer status. A full breakdown of each variable used in the study as shown in Table 1. Patients who underwent surgical resection were also analyzed for the amount of time from admission to surgery.

The above covariates were utilized to evaluate the effect of COVID-19 on the primary outcomes of (1) rate of surgical resection, (2) in-hospital mortality, and (3) discharge disposition. APR-DRG risk of mortality and disease severity was assigned using software developed by 3M Health Information Systems, and the ECI was grouped into four categories (0, 1, 2, and  $\geq$ 3 factors) based on an ICD-10 coding algorithm.

#### Statistical analysis

Descriptive analyses of patient demographics and clinical characteristics were stratified by COVID-19 status using mean (standard deviation) and *t*-test for continuous variables and frequency (percentage) and  $\chi^2$  tests for categorical variables. Two types of multivariable regression were then utilized to analyze outcomes: (1) logistic regression for surgical resection and in-hospital mortality

and (2) log-binomial regression for discharge disposition and transfer-in status. Generalized estimation equations were used in all models to account for hospital clustering. The logistic model assumptions were checked using the Hosmer-Lemeshow goodness-of-fit test. All of the covariates mentioned earlier were included in the model for adjustment. Age was recategorized as it violated the linearity assumption. Only complete cases were used for regressions without imputation since missing data affected <4% of records. P < 0.05 was considered statistically significant. The analysis was performed using SAS, version 9.4 (SAS Institute, Cary, North Carolina, USA).

## RESULTS

#### Patient demographics

A total of 30,671 patients met inclusion and exclusion criteria. Within this group, 738 (2.4%) patients had a COVID-19 diagnosis. A flowchart outlining the inclusion and exclusion process is shown in Figure 1. General characteristics of COVID-19-positive patients include older age (age 75+, 22.6% vs. 18.1%), higher APR-DRG risk of mortality (extreme, 48.6% vs. 25.3%), and severity of illness (extreme, 49.1% vs. 22.0%) [Table 1]. Additional patient characteristics pertaining to hospital stay are shown in Table 2.

#### Surgical resection cohort

COVID-19 patients who underwent surgical resection for malignant brain tumors trended toward a longer time from admission to surgery (3.1 days for COVID-19 patients vs. 2.1 days for patients without COVID-19); however, this finding was not statistically significant (P = 0.17). More COVID-19 patients also had 3 or more days from admission to surgery compared to patients without COVID-19 infection ( $\geq$ 3, 42.9% vs. 39.0%), but this difference was also not statistically significant (P = 0.25).

#### Patient outcomes

After controlling for other covariates, COVID-19 infection was significantly associated with a decreased likelihood of receiving surgery (Odds ratio [OR] 0.43, 95% confidence interval [CI] 0.29–0.63, P < 0.0001) [Table 2], an increased likelihood of mortality (OR 2.18, 95% CI 1.78–2.66, P < 0.0001) [Table 3], and increased likelihood of nonroutine discharge (OR 1.25, 95% CI 1.13–1.39, P < 0.0001) [Table 4].

Sub-cohort analysis of malignant primary tumors and malignant secondary (metastatic) tumors demonstrated similar results. In patients with malignant primary tumors, COVID-19 infection was associated with a decreased likelihood of receiving surgery (OR 0.37, 95% CI 0.22–0.62,

	COVID (-) <i>n</i> (%)	COVID (+) <i>n</i> (%)	Total cohort <i>n</i> (%)	P-value
Total patient count	29,933	738	30,671	
Age, years				0.0066^
18-44	2,966 (9.9)	57 (7.7)	3,023 (9.9)	
45-59	7439 (24.9)	175 (23.7)	7614 (24.8)	
60-74	14,113 (47.1)	339 (45.9)	14,452 (47.1)	
75+	5,415 (18.1)	167 (22.6)	5,582 (18.2)	
Gender				0.8295^
Male	14,837 (49.6)	358 (48.5)	15,195 (49.5)	
Female	15,094 (50.4)	380 (51.5)	15,474 (50.5)	
Missing	DS*	0 (0.0)	DS*	
Race				< 0.0001
Caucasian	21,493 (71.8)	439 (59.5)	21,932 (71.5)	
African-American	3,595 (12.0)	108 (14.6)	3,703 (12.1)	
Hispanic	2,157 (7.2)	107 (14.5)	2,264 (7.4)	
Asian or pacific islander	1,048 (3.5)	33 (4.5)	1,081 (3.5)	
Other	969 (3.2)	31 (4.2)	1,000 (3.3)	
Missing	671 (2.2)	20 (2.7)	691 (2.3)	
Surgery				< 0.0001
No	26,421 (88.3)	710 (96.2)	27,131 (88.5)	
Yes	3,512 (11.7)	28 (3.8)	3,540 (11.5)	
APR-DRG Severity of illness				< 0.0001
Minor	1,129 (3.8)	0 (0.0)	1,129 (3.7)	
Moderate	8,255 (27.6)	DS*	8,263 (26.9)	
Major	13,969 (46.7)	368 (49.9)	14,337 (46.7)	
Extreme	6,579 (22.0)	362 (49.1)	6,941 (22.6)	
APR-DRG risk of mortality				< 0.0001
Minor	1,118 (3.7)	0 (0.0)	1,118 (3.6)	
Moderate	6,685 (22.3)	DS	6,689 (21.8)	
Major	14,549 (48.6)	375 (50.8)	14,924 (48.7)	
Extreme	7,580 (25.3)	359 (48.6)	7,939 (25.9)	
Elixhauser				0.0434^
1	1,039 (3.5)	19 (2.6)	1,058 (3.4)	
2	2,994 (10.0)	57 (7.7)	3,051 (9.9)	
≥3	25,900 (86.5)	662 (89.7)	26,562 (86.6)	

Table 1: General patient demographics of all malignant brain tumor patients that met inclusion and exclusion criteria, stratified by COVID-19 diagnosis.

<sup>^</sup>Chi-square *P* value; <sup>^</sup>Equal variance two sample *t*-test, \*DS: Data suppressed for patient counts below 10 to protect patient anonymity per national inpatient sample policy. APR-DRG: All patient refined diagnosis related groups, COVID: Coronavirus disease, *n*: Number

P < 0.0002), increased likelihood of mortality (OR 1.91, 95% CI 1.15–3.16, P = 0.01), and increased likelihood of nonroutine discharge (OR 1.19, 95% CI 1.01–1.40, P = 0.04). In patients with malignant secondary tumors, COVID-19 infection was associated with a decreased likelihood of receiving surgery (OR 0.37, 95% CI 0.20–0.68, P = 0.001), increased likelihood of mortality (OR 2.36, 95% CI 1.88– 2.96, P < 0.0001), and increased likelihood of non-routine discharge (OR 1.24, 95% CI 1.08–1.43, P = 0.002).

Additional characteristics associated with likelihood of receiving surgery are shown in Table 3. The characteristics associated with a decreased likelihood of receiving surgery were older age (60–74-years-old: OR 0.76, 95% CI 0.67–0.86, P < 0.0001; 75+ years old: OR 0.57, 95% CI 0.48–0.67,

*P* < 0.0001) and ECI ≥3 (OR 0.38, 95% CI 0.32–0.44, *P* < 0.0001). Factors associated with increased likelihood of receiving surgery included private insurance (OR 1.38, 95% CI 1.22–1.57, *P* < 0.0001) and median income in the highest quartile (OR 1.21, 95% CI 1.06–1.38, *P* < 0.0001).

The factors associated with increased likelihood of mortality are shown in Table 4. Other than COVID-19 infection, characteristics associated with increased likelihood of mortality include older age (75+ years old: OR 1.28, 95% CI 1.03–1.60, P < 0.03) and non-white race.

Factors associated with likelihood of non-routine discharge are shown in Table 5. These included higher APR-DRG disease severity, ECI  $\geq$ 3 (OR 1.25, 95% CI 1.07–1.47, P = 0.005), and age over 45 years old.

	COVID (-) <i>n</i> (%)	COVID (+) <i>n</i> (%)	Total cohort n (%)	P-value
Total patient count	29,933	738	30,671	
Income				0.0219^
0–25 <sup>th</sup> percentile	7,367 (24.6)	220 (29.8)	7,587 (24.7)	
26 <sup>th</sup> –50 <sup>th</sup> percentile	7,791 (26.0)	172 (23.3)	7,963 (26.0)	
51 <sup>st</sup> –75 <sup>th</sup> percentile	7,241 (24.2)	176 (23.8)	7,417 (24.2)	
76 <sup>th</sup> -100 <sup>th</sup> percentile	7,149 (23.9)	160 (21.7)	7,309 (23.8)	
Missing	385 (1.3)	10 (1.4)	395 (1.3)	
Payer				< 0.0001
Medicare	14,853 (49.6)	409 (55.4)	15,262 (49.8)	
Medicaid	3,862 (12.9)	119 (16.1)	3,981 (13.0)	
Private insurance	9,500 (31.7)	171 (23.2)	9,671 (31.5)	
Self-pay	722 (2.4)	15 (2.0)	737 (2.4)	
Other	954 (3.2)	22 (3.0)	976 (3.2)	
Missing	42 (0.1)	DS*	44 (0.1)	
Hospital bed size				0.0006^
Small	4,422 (14.8)	144 (19.5)	4,566 (14.9)	
Medium	7,237 (24.2)	185 (25.1)	7,422 (24.2)	
Large	18,274 (61.0)	409 (55.4)	18,683 (60.9)	
Transfer-in				0.5527^
Missing	138 (0.5)	DS*	140 (0.5)	
No	24,224 (80.9)	590 (79.9)	24,814 (80.9)	
Yes	5,571 (18.6)	146 (19.8)	5,717 (18.6)	
LOS				< 0.0001^
N (missing)	29,933 (0)	738 (0)	30,671 (0)	
Mean (SD)	6.6 (6.93)	9.6 (10.11)	6.7 (7.03)	
In hospital mortality				< 0.0001
No	27,743 (92.7)	578 (78.3)	28,321 (92.3)	
Yes	2,190 (7.3)	160 (21.7)	2,350 (7.7)	
Discharge disposition	· · · ·	× •		< 0.0001
Unfavorable	9,160 (30.6)	388 (52.6)	9,548 (31.1)	
Favorable	20,773 (69.4)	350 (47.4)	21,123 (68.9)	

 Table 2: Additional patient demographics pertaining to hospital stay of all malignant brain tumor patients that met inclusion and exclusion criteria, stratified by COVID-19 diagnosis.

<sup>^</sup>Chi-square *P* value; <sup>^</sup>Equal variance two sample *t*-test, \*DS: Data suppressed for patient counts below 10 to protect patient anonymity per national inpatient sample policy. LOS: Length of stay, SD: **S**tandard deviation, COVID-19: Coronavirus disease 19, *n*: Number

# DISCUSSION

In our analysis of the NIS, malignant brain tumor patients with COVID-19 were associated with a decreased likelihood of receiving surgery, increased likelihood of in-hospital mortality, and increased likelihood of non-routine discharge. Notably, despite a decreased likelihood of surgery among COVID patients, those COVID patients receiving surgery did not demonstrate an association with surgical delay. To date, this study is the first to identify these findings using a large, national dataset and the findings of our study provide important insight into optimal management of these patients in resource-scarce circumstances like the COVID-19 pandemic.

Surgical resection is an essential component in the treatment of many malignant brain tumors. However, during the COVID-19 pandemic, many hospitals had to postpone or cancel non-emergent surgeries due to concerns about the risk of infection and need to preserve hospital resources. While often necessary, malignant tumor resections are typically not "emergent" and, as such, were often classified as "elective." However, due to the aggressive nature of malignant brain tumors, delaying surgery up to 7 weeks - as prior guidelines have recommended - is not always feasible.<sup>[17]</sup> Fortunately, we did note that COVID diagnosis was not necessarily associated with delay in surgery when patients were offered surgical resection. Further, prior limited studies have shown an association between COVID-19 and malignant brain tumors with poorer outcomes.<sup>[16]</sup> One study from the UK with 1221 patients found that the pandemic changed malignant brain tumor management, most commonly resulting in patients receiving no surgery or no treatment at all.<sup>[24]</sup> However, this study is limited by small sample size and limited generalizability and does not compare outcome

resection for malignant brain tumor patients.			
	Odds ratio (95% CI)	P-value	
COVID			
No	Reference		
Yes	0.43 (0.29-0.63)	< 0.0001	
Age			
<45	Reference		
45-59	0.84 (0.75-0.95)	0.0067	
60-74	0.76 (0.67-0.86)	< 0.0001	
75	0.57 (0.48-0.67)	< 0.0001	
Gender			
Male	Reference		
Female	0.79 (0.73-0.85)	< 0.0001	
Race			
Caucasian	Reference		
African-American	0.82 (0.72-0.94)	0.0041	
Hispanic	0.90 (0.77-1.06)	0.22	
Asian or Pacific Islander	0.63 (0.50-0.79)	< 0.0001	
Other	1.03 (0.84-1.26)	0.78	
Elixhauser comorbidity score			
1	Reference		
2	0.76 (0.64-0.90)	0.002	
≥3	0.38 (0.32-0.44)	< 0.0001	
APR-DRG disease severity			
Minor	Reference		
Moderate	2.60 (2.05-3.30)	< 0.0001	
Major	1.52 (1.20-1.93)	0.0006	
Extreme	1.06 (0.83-1.36)	0.65	
Bed size of hospital			
Large	Reference		
Medium	0.61 (0.53-0.70)	< 0.0001	
Small	0.44 (0.33-0.58)	< 0.0001	
Primary insurance			
Medicaid	Reference		
Medicare	1.14 (1.00-1.30)	0.048	
Private insurance	1.38 (1.22-1.57)	< 0.0001	
Self-pay	1.27 (1.00-1.62)	0.053	
Other	1.13 (0.89–1.44)	0.31	
Median household income			
0–25 percentile	Reference		
26–50 percentile	1.08 (0.97-1.20)	0.18	
51–75 percentile	1.09 (0.96-1.23)	0.18	
76–100 percentile	1.21 (1.06–1.38)	0.0046	
APR-DRG: All patient refined dia interval, Bolded: Significant value		onfidence	

Table 3: Factors associated with likelihood of receiving surgical

interval, Bolded: Significant values

between COVID-19-positive patients and COVID-19negative patients. Nevertheless, our study with a large, national dataset reinforces the findings from this prior study with a larger sample size.

In addition, our study found that patients with concurrent COVID-19 infection had 2.18 times greater odds of mortality than those without COVID-19. Patients with both cancer and COVID-19 infection pose a unique clinical

Reference 2.18 (1.78–2.66) Reference 0.98 (0.81–1.19) 1.14 (0.94–1.39) 1.28 (1.03–1.60) Reference 1.11 (1.01–1.22) Reference 1.30 (1.12–1.50) 1.22 (1.03–1.44) 1.43 (1.12–1.82) 1.17 (0.90–1.52)	<0.0001 0.83 0.19 0.029 0.031 0.0004 0.024 0.0038
2.18 (1.78–2.66) Reference 0.98 (0.81–1.19) 1.14 (0.94–1.39) 1.28 (1.03–1.60) Reference 1.11 (1.01–1.22) Reference 1.30 (1.12–1.50) 1.22 (1.03–1.44) 1.43 (1.12–1.82)	0.83 0.19 <b>0.029</b> <b>0.031</b> <b>0.0004</b> <b>0.024</b>
Reference 0.98 (0.81–1.19) 1.14 (0.94–1.39) 1.28 (1.03–1.60) Reference 1.11 (1.01–1.22) Reference 1.30 (1.12–1.50) 1.22 (1.03–1.44) 1.43 (1.12–1.82)	0.83 0.19 <b>0.029</b> <b>0.031</b> <b>0.0004</b> <b>0.024</b>
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0.98 (0.81–1.19) 1.14 (0.94–1.39) 1.28 (1.03–1.60) Reference 1.11 (1.01–1.22) Reference 1.30 (1.12–1.50) 1.22 (1.03–1.44) 1.43 (1.12–1.82)	0.19 0.029 0.031 0.0004 0.024
1.14 (0.94–1.39) 1.28 (1.03–1.60) Reference 1.11 (1.01–1.22) Reference 1.30 (1.12–1.50) 1.22 (1.03–1.44) 1.43 (1.12–1.82)	0.19 0.029 0.031 0.0004 0.024
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1.30 (1.12–1.50) 1.22 (1.03–1.44) 1.43 (1.12–1.82)	0.024
1.22 (1.03–1.44) 1.43 (1.12–1.82)	0.024
1.43 (1.12–1.82)	
	0.0038
1.17 (0.90-1.52)	
	0.24
Reference	
1.128 (0.719–1.767)	0.60
1.407 (0.949–2.088)	0.090
Reference	
2.07 (0.90-4.79)	0.090
5.99 (2.39–14.98)	0.0001
38.70 (15.05–99.50)	<0.0001
Reference	
	0.0017
1.402 (1.113–1.765)	0.0041
Reference	
	0.0005
0.90 (0.78-1.05)	0.19
1.27 (0.93–1.73)	0.14
1.91 (1.46–2.49)	<0.0001
Reference	
	0.092
	0.31
0.90 (0.77-1.05)	0.17
	1.128 (0.719–1.767) 1.407 (0.949–2.088) Reference 2.07 (0.90–4.79) 5.99 (2.39–14.98) 38.70 (15.05–99.50) Reference 1.218 (1.077–1.378) 1.402 (1.113–1.765) Reference 0.76 (0.65–0.89) 0.90 (0.78–1.05) 1.27 (0.93–1.73) 1.91 (1.46–2.49)

dilemma – in the interest of not advancing their COVID-19 infection, it may be sensible to withhold chemotherapy and steroid treatment so as to not further immunosuppress these patients.<sup>[34]</sup> In addition, COVID-19 treatments themselves may immunosuppress patients allowing for metastatic progression.<sup>[29]</sup> The largest study to date of 869 neurosurgical and neurointerventional patients at two medical centers demonstrated that those with a concurrent COVID-19 diagnosis are associated with an increased risk

discharge for malignant brain	n tumor patients.	
	Odds ratio (95% CI)	P-value
COVID		
No	Reference	
Yes	1.24 (1.16-1.33)	<0.0001
Age		
<45	Reference	
45-59	1.23 (1.14-1.33)	<0.0001
60-74	1.39 (1.29–1.49)	< 0.0001
≥75	1.59 (1.46–1.72)	< 0.0001
Gender		
Male	Reference	
Female	0.98 (0.95-1.01)	0.20
Race		
Caucasian	Reference	
African-American	1.07 (1.02–1.12)	0.0089
Hispanic	0.93 (0.87-1.00)	0.062
Asian or Pacific Islander	0.90 (0.82-0.99)	0.033
Other	1.03 (0.93–1.13)	0.61
Elixhauser comorbidity score		
1	Reference	
2	1.06 (0.91–1.23)	0.44
≥3	1.26 (1.10–1.44)	0.0010
APRDRG disease severity		
Minor	Reference	
Moderate	1.37 (1.17–1.61)	0.0001
Major	1.97 (1.68–2.31)	< 0.0001
Extreme	3.48 (2.97-4.09)	< 0.0001
Bed size of hospital		
Large	Reference	
Medium	1.07 (1.02–1.12)	0.012
Small	1.08 (1.02–1.15)	0.012
Primary insurance		
Medicare	Reference	
Medicaid	1.03 (0.97-1.09)	0.32
Private insurance	0.87 (0.83-0.92)	< 0.0001
Self-pay	0.76 (0.66–0.87)	0.0001
Other	1.09 (1.00–1.19)	0.051
Median household income	· · · · /	
0–25 percentile	Reference	
26–50 percentile	0.99 (0.94-1.03)	0.53
51–75 percentile	1.00 (0.95–1.05)	0.87
76–100 percentile	0.97 (0.91–1.03)	0.27
APR-DRG: All patient refined di interval, Bolded: Significant valu	iagnosis related groups, CI: C	onfidence

Table 5: Factors associated with likelihood of unfavorable

discharge for malignant brain tumor patients

interval, Bolded: Significant values

of mortality (18.2% vs. 4.1%, P = 0.0029).<sup>[28]</sup> The same study also found COVID-19 to be associated with non-routine discharge, which was similarly identified in our study. One international meta-analysis found that the pooled mortality risk estimate for hospitalized patients with COVID-19 and cancer was 14.1%, at least 5 times greater than the mortality risk of non-elderly patients with COVID-19.<sup>[38]</sup> Furthermore, a small study of patients with primary brain tumors and COVID-19 in the Netherlands found that half of the patients had a severe disease course requiring hospital admission and 13% had a fatal outcome.<sup>[9]</sup> A recent cohort study in China found that patients with brain cancer are at increased risk for COVID-19 mortality.<sup>[6]</sup>

Further analysis is necessary to elucidate the etiology of increased mortality among patients with brain tumors who have COVID-19. Intrinsic medical factors, such as cytokine storm, neuroinflammation, and endothelial cell destruction secondary to COVID-19 infection, could contribute to worse outcomes in brain tumor patients.<sup>[10,25]</sup> Other logistical contributors to worse outcomes could be delays in cancer treatment, immunosuppression from chemotherapy, and limited hospital resources.<sup>[12]</sup>

Our analysis also found demographic differences between patients with malignant brain tumors with and without COVID-19. A higher percentage of malignant brain tumor patients with COVID-19 were African-American (14.9%) and Hispanic (14.5%) compared to patients without COVID-19 (12% were African-American and 7.2% were Hispanic). Racial disparities in COVID-19 infections and outcomes have been well documented in the United States; African-American and Hispanic populations have an estimated 1.5-3.5 higher risk and 1.3-7.7 higher risk, respectively, compared to white populations.<sup>[14,15,33]</sup> Our data further validate these trends among malignant brain tumor patients in the United States. In addition, we found that malignant brain tumor patients with COVID-19 were significantly more likely to have Medicare or Medicaid rather than private insurance compared to malignant brain tumor patients without COVID-19. Insurance status has been linked with hospitalization rates from COVID-19; one 2020 retrospective and cohort study found that patients with Medicaid were more than 2 times as likely to admitted to a hospital for COVID-19 than those with commercial insurance.<sup>[4]</sup> One retrospective and cohort study conducted in Georgia found that what was assumed to be a racial disparity in COVID-19 mortality can be attributed to insurance status; after controlling for insurance status, African-American and Caucasian patients had similar mortality outcomes, but patients with private insurance had significantly lower mortality rates than those with Medicare (OR = 0.68, 95% CI 0.48-0.96).<sup>[17]</sup> Further analysis is needed to determine the impact of race and insurance status on outcomes among patients with malignant brain tumors with and without COVID-19.

#### Limitations

Our retrospective and cohort study, which utilized the NIS, has several inherent limitations. Although the NIS contains a large number of patients, only 1 year of 2020 data has been made public so far, limiting the number of documented COVID-19 patients who underwent tumor resection. The analysis was therefore underpowered, which may affect the statistical significance noted in assessing the association of COVID with time to surgery. The NIS also cannot identify patients with multiple hospitalizations, resulting in readmission being recorded as new patients. This limits the ability to track patients with long-term conditions and assess for long-term cancer survival. Furthermore, the NIS does not distinguish pre-existing conditions and hospital-acquired conditions. In addition, the NIS is specific to the United States, and since the impact of the COVID-19 pandemic varies greatly across different regions and countries, NIS data may not be generalizable. Finally, as with any administrative database, reliance on ICD-10 coding can result in errors and subjectivity in data input and coding. Despite these limitations, our study provides valuable insights into the relationship between malignant brain tumors and COVID-19. Further research is needed to validate our findings and to address the limitations of the NIS.

# CONCLUSION

Overall, malignant brain tumor patients with COVID-19 infection were associated with worse overall outcomes, including decreased likelihood of receiving surgery, increased likelihood of mortality, and increased likelihood of non-routine discharge. To the best of our knowledge, this study is the first to analyze the relationship of these outcomes with COVID-19 infection with a large national dataset. The findings of our study emphasize the need to balance the risks and benefits of delaying surgery for malignant brain tumor patients with COVID-19. Further study on causal relationships between COVID-19 and patient outcomes is needed to optimize management for this patient population at risk for direct effects of COVID infection and indirect health-care system consequences of COVID-19. As such, though the COVID-19 pandemic is no longer a public health emergency, reinforcing our understanding of the pandemic's impact on outcome provides important insight in effective triage for these patients in future situations where resources are limited.

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# Declaration of patient consent

Patients' consent not required as patients' identities were not disclosed or compromised.

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## **Conflicts of interest**

There are no conflicts of interest.

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# SUPPLEMENTARY TABLE

Supplementary Table S1: ICD-10 Codes u	used in inclusion criteria.			
Disease category: Malignant				
Diagnosis code description ICD-10	Diagnosis code ICD-10	Procedure code description ICD-10	Procedure code ICD-10	
Malignant neoplasm of the brain	C71.0-71.9	Excision of Cerebral Hemisphere – Open, Percutaneous, and Percutaneous endoscopic approach Destruction/Excision of Brain – Open, Percutaneous, and Percutaneous Endoscopic Approach	00B70ZZ, 00B73ZZ, and 00B74ZZ 00500ZZ, 00503ZZ, 00504ZZ, 00B00ZZ, 00B03ZZ, and 00B04ZZ	
Secondary malignant neoplasm of the brain	C79.31	Destruction/Excision of Brain – Open, Percutaneous, and Percutaneous Endoscopic Approach	00500ZZ, 00503ZZ, 00504ZZ, 00B00ZZ, 00B03ZZ, and 00B04ZZ	
Malignant neoplasm of unspecified cranial nerve, malignant neoplasm of other cranial nerves	C72.50 and C72.59	Excision of Cerebral Hemisphere – Open, Percutaneous, and Percutaneous Endoscopic Approach	00B70ZZ, 00B73ZZ, and 00B74ZZ	
Malignant neoplasm of central nervous system, unspecified	C72.9			
Secondary malignant neoplasm of brain	C79.31			
Secondary malignant neoplasm of	C79.32, C79.49, and			
cerebral meninges, secondary malignant neoplasm of unspecified part of nervous system, secondary malignant neoplasm of other parts of nervous system	C79.40			
Malignant neoplasm of cerebral meninges; malignant neoplasm of meninges, unspecified	C70.0 and C70.9	Destruction/Excision of Cerebral Meninges – Open, Percutaneous, and Percutaneous Endoscopic Approach	00510ZZ, 00513ZZ, 00514ZZ, 00B10ZZ, 00B13ZZ, and 00B14ZZ	

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